

Lecture 1 - Probability Basics

①

A student in ACLC hall is randomly selected and the student is from COE: College of Engineering, COSAM: College of Sciences and Math or oth: other

Let X = college of the student

$X = \text{COE}$ outcome ω
 $X = \text{COSAM}$
 $X = \text{oth}$

Sample space

$$\Omega = \{\text{COE}, \text{COSAM}, \text{oth}\}$$

- An event, A , is any subset of Ω : $A \subseteq \Omega$
- Def'n: Probability P : subsets of $\Omega \rightarrow \mathbb{R}$, actually $[0, 1]$
- Suppose
 $P(X = \text{COE}) = 0.5$, $P(X = \text{COSAM}) = 0.3$,
 $P(X = \text{oth}) = 0.2$

$$\sum_{i \in \Omega} P(X=i) = 1 \quad \checkmark$$

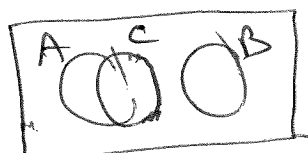
If $X = \text{COE}$, then $P(X = \text{COSAM}) = 0$

Mutually exclusive (ME) Events:

$E_1, E_2, \dots, E_n$ are mutually exclusive if
 $E_i \cap E_j = \emptyset$ for all $i \neq j$

Note:

① $A \cap B = \emptyset$ and $B \cap C = \emptyset$, $A \cap C \neq \emptyset$



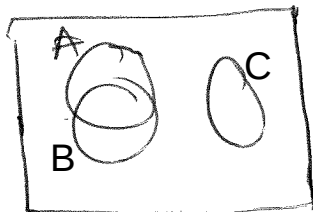
"Is $A \cap C = \emptyset$?"

Not necessarily

← see

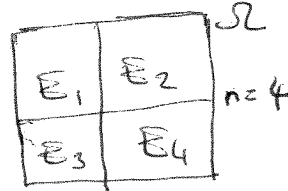
(2)

② If $A \cap B \cap C = \emptyset$, then $A \cap B = \emptyset$
Not necessarily!



• Def'n $E_1, \dots, E_n$ is a partition of Ω or
" partition Ω ,

if (i) events are mutually exclusive
and (ii) $\bigcup_{i=1}^n E_i = \Omega$ $E_1 \cup E_2 \cup \dots \cup E_n = \Omega$



Ex 1: $\{C \cap E\}, \{C \cap A \cap M\}, \{\text{oth}\}$ is a partition (of Ω)
 $\{C \cap E, \text{oth}\}, \{C \cap A \cap M\}$ " another " " " ,

Assigning Probabilities

Method 1: Set Theory

$$P(X \in \{C \cap E \text{ or } C \cap A \cap M\}) = 0.8$$

$$P(C \cap E \cup C \cap A \cap M) = P(C \cap E) + P(C \cap A \cap M) - P(C \cap E \cap C \cap A \cap M)$$

(inclusion-exclusion)

$$\text{In general } = 0.5 + 0.3 - 0 = 0.8.$$



ME Rule: ① If E_1 & E_2 are ME, then

$$P(E_1 \cup E_2) = P(E_1) + P(E_2)$$

② If $E_1, E_2, \dots, E_n$ are ME, then

$$P\left(\bigcup_{i=1}^n E_i\right) = \sum_{i=1}^n P(E_i)$$

Def'n P : Subsets of $\Omega \rightarrow \mathbb{R}$

(i) $P(A) \geq 0$ for all $A \subseteq \Omega$

(ii) $P(\Omega) = 1$

(iii) $P(\bigcup_{i=1}^{\infty} E_i) = \sum_{i=1}^{\infty} P(E_i)$ for ME events $E_i \subseteq \Omega$

Applications: (i) $P(A^c) = 1 - P(A)$

Sol'n: $P(A \cup A^c) = P(\Omega) = 1$

Since A & A^c are ME, $P(A \cup A^c) = P(A) + P(A^c)$

So, $P(A) + P(A^c) = 1 \Rightarrow P(A^c) = 1 - P(A)$.

(2) Union Bound: $P(A \cup B) \leq P(A) + P(B)$ for any events A, B

Sol'n: $P(A \cup B) = P(A) + P(B) - P(A \cap B)$

Since $P(A \cap B) \geq 0$, we have

$$P(A) + P(B) - P(A \cap B) \leq P(A) + P(B)$$

$$\Rightarrow P(A \cup B) \leq P(A) + P(B)$$

In general, $P(A \cup B \cup \dots) \leq P(A) + P(A) + \dots$

Def'n: A and B are independent iff $P(A \cap B) = P(A)P(B)$
 $A \perp B$

(1) A, B, C are pairwise indep. if $A \perp B, A \perp C, B \perp C$

(2) A, B, C " mutually " if $P(A \cap B \cap C) = P(A)P(B)P(C)$

Method 2 Counting (works for equally likely outcomes)

Assume a student's getting A, B or C are equally likely.

$$P(\text{grade better than } C) = 2/3$$

(4)

In general, $P(E) = \frac{\# \text{ of outcomes favorable to } E}{|\Omega|}$

$$= \frac{|E|}{|\Omega|}$$

X & Y are two students & get grades A, B, C in pairs equally likely
 $P(X=A \cap Y=A) \stackrel{?}{=} P(X=A)P(Y=A)$
 $\{A, A\}, \{A, B\}, \dots \quad \frac{1}{9} \quad \frac{1}{3} \times \frac{1}{3} \quad \checkmark \quad \text{indep.}$

Ex. $\{A, B, C\}$ equally likely, 10 students
 $P(\text{exactly 3 students get } A)$

$$= \frac{\binom{10}{3} 2^7}{3^{10}}$$

$$|\Omega| = 3^{10}$$

$\underbrace{A \ A \ A}_{3 \text{ get } A} \underbrace{B/C \ B/C \ \dots \ B/C}_{2^7}$

$\{A, A, \dots\}$

$\{A, A, A, B, \dots\}$

$\{A, A, B, B, \dots, C, B, A\}$

$$= \binom{10}{3} \left(\frac{1}{3}\right)^3 \left(\frac{2}{3}\right)^7 \quad \checkmark$$

Extension 1: $P(\text{at most 3 stds get } A)$

$$= \binom{10}{0} \left(\frac{1}{3}\right)^0 \left(\frac{2}{3}\right)^{10} + \binom{10}{1} \left(\frac{1}{3}\right)^1 \left(\frac{2}{3}\right)^9 + \binom{10}{2} \left(\frac{1}{3}\right)^2 \left(\frac{2}{3}\right)^8 + \binom{10}{3} \left(\frac{1}{3}\right)^3 \left(\frac{2}{3}\right)^7$$

Extension 2: $P(\text{at least 4 stds get } A)$

$$= 1 - P(\text{at most 3 stds get } A)$$

Qu: Suppose $A \perp B$. Can A & B be ME?

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Ans: $P(A \cap B) = P(A)P(B)$ since $A \perp B$

For A & B to be ME, $A \cap B = \emptyset$ must hold, which would imply $P(A \cap B) = 0$

So, $P(A)P(B) = 0$ iff $P(A) = 0$ or $P(B) = 0$.

Thus A & B can be ME, if $A = \emptyset$ or $B = \emptyset$

Def'n $RND(a, b)$ generates all integers between a & b (inclusive) with equal prob.

$$RND(1, 3) = \begin{cases} 1 & \text{wp } 1/3 \\ 2 & \text{" " " } \\ 3 & \text{" " " } \end{cases}$$

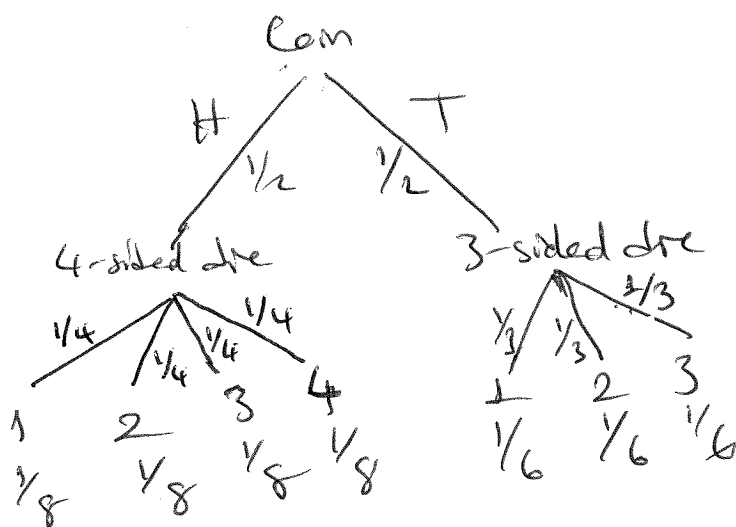
die: $RND(1, 6)$

com: $RND(0, 1)$

Method 3 Trees

Ex: 1. toss a fair coin

2. If coin shows Heads, roll a 4-sided fair die
else " " 3-sided " " "



① $P(\text{getting a 4 m 4-sided die}) = ?$

$$\frac{1}{2} \times \frac{1}{4} = \frac{1}{8}$$

② $P(\text{die} \geq 3)$

$$= \frac{1}{8} + \frac{1}{8} + \frac{1}{6} = \frac{5}{12}$$

Notes: (1) Branches form a partition

(2) Could use $RND(0, 1)$

$RND(1, 3)$

$RND(1, 4)$

(6)

Ex 1 We have 100 indep. tosses of an unbiased coin.

If 99 tosses result in a H, what is the prob
 $P(\text{100th toss} = H) = ? \quad \frac{1}{2}$

Ex 2: Roll two unbiased dice, D_1 & D_2 , independently,

(1) Is $D_1 + D_2 = 5 \perp D_2 = 6$? No,

(2) " " " " $D_2 = 5$? No

(3) " " " " $D_2 = 4$? No

(4) " " " " $D_2 = y$: $y \in \{1, 2, \dots, 6\}$ Yes

(5) " " $= 3 \perp D_2 = y$ " No

(6) " " $= 7 \perp D_2 = 3$ Yes!

$$\underbrace{P(D_1 + D_2 = 7 \cap D_2 = 3)}_{1/36} \stackrel{?}{=} \underbrace{P(D_1 + D_2 = 7)}_{\frac{6}{36} = 1/6} \times \underbrace{P(D_2 = 3)}_{1/6}$$

Qui: $D_1 + D_2 = 20 \perp D_2 = 4$ Yes!

• For any event A , $A \perp \emptyset$

Sol'n: $P(A \cap \emptyset) \stackrel{?}{=} P(A) P(\emptyset)$

$$P(\emptyset) \stackrel{?}{=} P(A) \cdot 0$$

$$0 = 0 \quad \checkmark$$

• $A \perp B \Rightarrow A \perp B^c$

" $\Rightarrow A^c \perp B$

" $\Rightarrow A^c \perp B^c$

⑦

Recall $\{CoE, CoSAM, oth\}$

Qn: 6 students are randomly selected; 2 from each of CoE, CoSAM, oth colleges.

(1) $P(\text{first std from CoE}) = P(S_1 = CoE) = ?$

$\{\underline{CoE}, \underline{CoE}, CoSAM, CoSAM, oth, oth\} \quad \frac{2}{6} = \frac{1}{3}$

or $|S| = \frac{6!}{2!2!2!} = 90 \quad \& \quad |S_1 = CoE| = \frac{5!}{2!2!} = 30$

$\Rightarrow P(S_1 = CoE) = \frac{30}{90} = \frac{1}{3}$

(2) $P(\text{first 3 stds from different colleges})$

$= P(S_{1-3} \in \{CoE, CoSAM, oth\}) = ?$

$\underbrace{CoE, CoSAM, oth}_{S_1, S_2, S_3}, \underbrace{CoE, CoSAM, oth}_{\text{remaining 3 stds}} = 36$

$\Rightarrow \frac{36}{90} = \frac{2}{5}$

or $P(S_{1-3} \in \{CoE, CoSAM, oth\})$

$= \frac{2}{6} \times \frac{2}{5} \times \frac{2}{4} \times 3! = \frac{2}{5}$

(3) $P(S_{1-3} \in \{CoE, CoSAM, oth\} \cap S_6 = oth)$

$= \frac{3! \cdot 2!}{90} = \frac{2}{15}$

$\underbrace{CoE, CoSAM, oth}_{S_1, S_2, S_3}, \underbrace{CoE, CoSAM}_{S_4, S_5}, \overset{\uparrow}{S_6}$

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or tree approach

$$P(S_{1-3} \in \{ \text{CoE, CoSAM, oth} \} \cap S_6 = \text{oth})$$

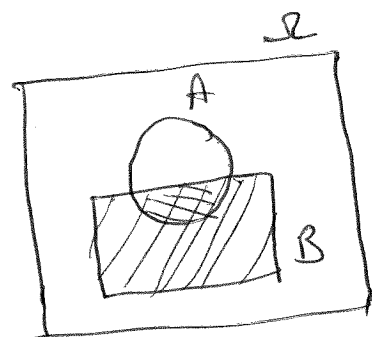
$$= P(S_{1-3} \in \{ \text{CoE, CoSAM, oth} \}) \times P(S_6 = \text{oth} \text{ given } S_{1-3} \in \{ \text{CoE, CoSAM, oth} \})$$

$$= \frac{2}{5} \times \frac{1}{3} = \frac{2}{15}$$

Conditional Probability

$$P(A|B) = \frac{P(A \cap B)}{P(B)}, \quad P(B) > 0$$

"prob of A given B"



$$\Rightarrow P(A \cap B) = P(A|B)P(B)$$

Similarly $P(B|A) = \frac{P(B \cap A)}{P(A)} \Rightarrow P(A \cap B) = P(B|A)P(A)$

• $P(A|B) \neq P(B|A)$ unless $P(A) = P(B)$

• If $A \perp B$, then $P(A \cap B) = P(A)P(B) = P(B)P(A|B)$
 $\Rightarrow P(A|B) = P(A)$

• $P(\cdot|B)$ is a valid prob. function (verify)

Chain Rule: $P(A \cap B) = P(A|B)P(B)$

$$P(A \cap B \cap C) = P(A|B \cap C)P(B \cap C)$$

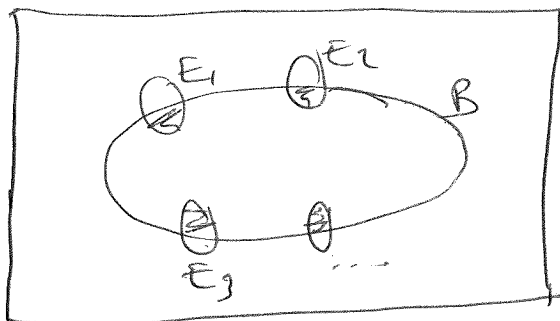
$$= P(A|B \cap C)P(B|C)P(C)$$

→ If $E_1, E_2, \dots$ are ME, rules of prob $\Rightarrow$

$$P(\cup E_i) = \sum P(A_i)$$

$$\Rightarrow P(\cup E_i | B) = \sum P(A_i | B)$$

(9)



Total Probability (TP) Rule:

If $\{E_1, E_2, \dots, E_n\}$ forms a partition of Ω and $A \subseteq \Omega$ is some event, then

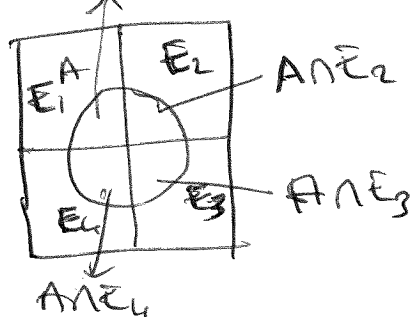
$$P(A) = \sum_{i=1}^n P(A|E_i)P(E_i)$$

Sol'n $P(A) = P(A \cap \Omega) = P(A \cap (\bigcup_{i=1}^n E_i))$

$$= P(\bigcup_{i=1}^n A \cap E_i)$$

$$\stackrel{\text{ME}}{=} \sum_{i=1}^n P(A \cap E_i) \stackrel{\substack{\uparrow \\ \text{chain rule}}}{=} \sum_{i=1}^n P(A|E_i)P(E_i)$$

with $n=4$



Bayes Thm:

(Simple Version)

Recall $P(A|B) = \frac{P(A \cap B)}{P(B)}$

$$\Rightarrow P(A \cap B) = P(A|B)P(B)$$

$$\& \quad P(B|A) = \frac{P(B \cap A)}{P(A)}$$

$$\Rightarrow P(A \cap B) = P(B|A)P(A)$$

$$\Rightarrow P(A|B)P(B) = P(B|A)P(A) \Rightarrow P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$

(general version) If $\{E_1, E_2, \dots, E_n\}$ forms a partition of Ω and A is any event, then (10)

$$P(E_i|A) = \frac{P(A|E_i)P(E_i)}{\sum_{j=1}^n P(A|E_j)P(E_j)}$$

Pf. $P(E_i|A) = \frac{P(A \cap E_i)}{P(A)} \xrightarrow{\text{chain rule}} \frac{P(A|E_i)P(E_i)}{\sum_{j=1}^n P(A|E_j)P(E_j)}$
 $\xrightarrow{\text{TP rule}}$

Ex: H : healthy
 D : diseased (not healthy), testing people for the disease

$$P(+|D) = 0.99 \Rightarrow P(-|D) = 0.01$$

$$P(-|D) = 0.99 \Rightarrow P(+|H) = 0.01$$

There is a $\frac{1}{1000}$ chance of catching the disease, i.e.
 $P(D) = \frac{1}{1000}$

Q: $P(D|+) = ?$

$$P(D|+) = \frac{P(+|D)P(D)}{P(+|D)P(D) + P(+|H)P(H)}$$

$$= \frac{0.99 \times \frac{1}{1000}}{0.99 \times \frac{1}{1000} + 0.01 \times \frac{999}{1000}} = \frac{0.99}{0.99 + 9.99} \approx \frac{1}{11}$$

Paradox: A couple has 2 children and at least one of them is a girl.

Q: Prob(both children are girls | at least 1 is a girl) = ?

Sol'n $P(\text{both are girls} | \geq 1 \text{ is a girl})$

$$= \frac{P(\text{both GG} | \geq 1 G)}{P(\geq 1 G)}$$

BB, BG, GB, GG

$$= \frac{P(GG)}{P(\geq 1 G)} = \frac{\frac{1}{4}}{\frac{3}{4}} = \frac{1}{3} \quad (\text{not } \frac{1}{2})$$

Conditional Independence:

Event A & B are $\perp$ conditional on the event C

$$\Leftrightarrow P(A \cap B | C) = P(A | C) P(B | C)$$

Notes:

(i) $\perp \not\Rightarrow$ Cond $\perp$.

Ex: $A: D_1 = 4$ $B: D_2 = 4$, $A \perp B$

But with $C: D_1 + D_2 = 6$, $A|C \not\perp B|C$

verify! $P(A \cap B | C) = \frac{P(A \cap B \cap C)}{P(C)} = \frac{P(\emptyset)}{P(C)} = 0$

but $P(A | C) = \frac{P(A \cap C)}{P(C)} = \frac{P(D_1 = 4, D_2 = 2)}{P(D_1 + D_2 = 6)} = \frac{1/36}{5/36} = \frac{1}{5}$

likewise $P(B | C) = 1/5$.
(1,5) (5,1), (2,4) (4,2), (3,3)

& $0 \neq \frac{1}{5} \times \frac{1}{5}$

(ii) Conditional $\perp \not\Rightarrow \perp$

Ex: $A|C: S_1$ carries a jacket if it snows w.p 0.999

$B|C: S_2$ " " " " " " 0.980

$C: It is snowing$

Assume $A|C \perp B|C$ i.e. $P(A \cap B | C) = P(A | C) P(B | C)$
 $= 0.999 \times 0.980$

But this has no implication on $P(A \cap B) = P(A)P(B)$!